



Markov approximation and the generalized entropy ergodic theorem for non-null stationary process

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Abstract. In an earlier work, we proved a generalized entropy ergodic theorem for finite nonhomogeneous Markov chains (NMC). In this paper, we establish a generalized strong law of large numbers for finite m -th order NMC. Then we deduce a generalized entropy ergodic theorem for finite m -th order NMC, under some assumptions on the continuity rate and of non-nullness. Explicit upper and lower bounds relating the generalized relative entropy density of the original finite non-null stationary sequence and its canonical m -order Markov approximation is obtained.

Keywords. m -th order nonhomogeneous Markov chains; non-null stationary process; canonical approximation.

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1. Introduction

Let $X = (X_n)_{n \in \mathbb{N}}$ be a discrete stochastic process taking values on a finite alphabet $\mathbf{X} = \{1, 2, \dots, b\}$ and defined on a probability space $(\Omega, \mathbf{F}, \mathbb{P})$. In the sequel, we use the convention that $\mathbb{N}^* = \mathbb{N} \setminus \{0\}$. Given two integers $m \leq n$, let X_m^n and x_m^n be the strings (X_m, \dots, X_n) and $(x_m, \dots, x_n) \in \mathbf{X}^{n-m+1}$ respectively. The subscript is omitted when it is 1. Given two strings $x^m = (x_1, \dots, x_m) \in \mathbf{X}^m$ and $y^n = (y_1, \dots, y_n) \in \mathbf{X}^n$, we denote their concatenation in \mathbf{X}^{m+n} by $x^m y^n$. Write

$$p(x_m^n) = \mathbb{P}(X_m^n = x_m^n), \quad x_k \in \mathbf{X}, \quad m \leq k \leq n$$

and, if $p(x_0^{m-1}) > 0$, we write

$$p(a|x_0^{m-1}) = \mathbb{P}(X_m = a | X_0^{m-1} = x_0^{m-1}).$$

For $m = 0$, $p(a|x_0^{m-1}) = p(a)$.

Let $(a_n, \phi(n))_{n \in \mathbb{N}}$ be a sequence of pairs of positive integers with $\phi(n)$ tending to infinity as $n \rightarrow \infty$. Set

$$f_{a_n, \phi(n)}(\omega) := -\frac{1}{\phi(n)} \log p(X_{a_n+1}^{a_n+\phi(n)}). \quad (1.1)$$

The function $f_{a_n, \phi(n)}(\omega)$ will be called the generalized relative entropy density of $X_{a_n+1}^{a_n+\phi(n)}$. In particular, if $a_n \equiv 0$ and $\phi(n) = n$, $f_{0,n}(\omega)$ denotes the classical relative entropy density of X^n , i.e.,

$$f_{0,n}(\omega) = -\frac{1}{n} \log p(X^n). \quad (1.2)$$

Hereafter, \log denotes the natural logarithm unless stated otherwise.

The convergence of $f_{0,n}(\omega)$ to a constant in the sense of \mathbf{L}_1 convergence, convergence in probability or a.e. convergence is called the Shannon–McMillan–Breiman theorem or the individual ergodic theorem of information or the asymptotic equipartition property (AEP) in information theory. There is a lot of research on this topic. Shannon [16] gave the original version for convergence in probability for stationary ergodic information sources with finite alphabet. McMillan [13] and Breiman [5, 6] obtained the entropy ergodic theorem in \mathbf{L}_1 and a.e. convergence, respectively, for finite stationary ergodic information sources. Chung [7] considered the case of countable alphabet. Billingsley [4] extended the result to stationary nonergodic sequences. The entropy ergodic theorem for general stochastic processes can be found, for example, in Barron [2], Kieffer [12] or Algoet and Cover [1]. Yang [18] obtained the entropy ergodic theorem for a class of nonhomogeneous Markov chains. Yang and Liu [19] proved the entropy ergodic theorem for a class of m -th order nonhomogeneous Markov chains and Zhong *et al.* [20] proved the entropy ergodic theorem for a class of asymptotic circular Markov chains.

In this paper, we will consider the convergence of $f_{a_n, \phi(n)}(\omega)$ and call it the generalized entropy ergodic theorem when $f_{a_n, \phi(n)}(\omega)$ converges to a constant in the sense of $\mathbb{P}|_{\sigma(X)}$ a.e. convergence. We should mention some recent contributions on this aspect. The first is the work of Nair [14], in which he established a moving average version of the Shannon–McMillan–Breiman theorem.

Theorem A. *Let $(X_n)_{n \in \mathbb{N}}$ be a two-sided stationary process taking values from the finite set $K = \{a_1, \dots, a_s\}$ and let $p(x_0, \dots, x_n)$ denote the joint distribution function of the variables X_0, \dots, X_n . If $(n_l, k_l)_{l \in \mathbb{N}^*}$ is of Stoltz [10], then there is a constant H such that*

$$\lim_l -\frac{1}{k_l} \log p(X_{n_l}^{n_l+k_l}) = H \quad \text{a.e.}$$

He gave an interesting illustration of this new theorem.

The second is by Wang and Yang [17], which proved that for a non-homogeneous Markov chain, the generalized relative entropy density $f_{a_n, \phi(n)}(\omega)$ converges a.e. and in \mathbf{L}_1 to the entropy rate of the Markov chain.

Theorem B. *Let $(X_n)_{n \in \mathbb{N}}$ be nonhomogeneous Markov chains with their transition probability matrices $P_n = (p_n(i, j))_{b \times b}$, $n \in \mathbb{N}^*$. Let $(a_n)_{n \in \mathbb{N}}$, $(\phi(n))_{n \in \mathbb{N}}$ be two sequences of nonnegative numbers such that, for every $\varepsilon > 0$, we have $\sum_{n=1}^{\infty} \exp[-\varepsilon \phi(n)] < \infty$. Let $P = (p(i, j))_{b \times b}$ be another irreducible transition matrix. If*

$$\lim_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} |p_k(i, j) - p(i, j)| = 0, \quad \forall i, j \in \mathbf{X},$$

then

$$\lim_n f_{a_n, \phi(n)}(\omega) = - \sum_{i=1}^b \sum_{j=1}^b \pi_i p(i, j) \log p(i, j) \quad \text{a.e. and in } \mathbf{L}^1,$$

where (π_1, \dots, π_b) is the unique stationary distribution determined by the transition matrix P .

Before going further, we first consider the notion of a *non-null* process.

DEFINITION 1

Let $X = (X_n)_{n \in \mathbb{N}}$ be a stationary stochastic process with state space \mathbf{X} . The process X is called *non-null* if, for any $k \geq 0$, we have $p(X_0^{k-1}) > 0$ and in addition,

$$p_{\text{inf}} := \inf_{k \geq 1} \min_{a \in \mathbf{X}, x_0^{k-1} \in \mathbf{X}^k} p(a|x_0^{k-1}) > 0. \tag{1.3}$$

The process X is continuous if

$$\beta(k) := \sup_{j \geq k} \max_{a \in \mathbf{X}} \max_{x_0^{j-1}, y_0^{j-1} \in \mathbf{X}^j: x_{j-k}^{j-1} = y_{j-k}^{j-1}} |p(a|x_0^{j-1}) - p(a|y_0^{j-1})| \rightarrow 0 \text{ as } k \rightarrow \infty. \tag{1.4}$$

The sequence $(\beta(k))_{k \in \mathbb{N}}$ is called the continuity rate.

Remark 1. A strong notion of continuity, often used in the literature [8], involves the log-continuity rate, namely

$$\gamma(k) := \sup_{j \geq k} \max_{a \in \mathbf{X}} \max_{x_0^{j-1}, y_0^{j-1} \in \mathbf{X}^j: x_{j-k}^{j-1} = y_{j-k}^{j-1}} \left| \frac{p(a|x_0^{j-1})}{p(a|y_0^{j-1})} - 1 \right|. \tag{1.5}$$

The process X is log-continuous if $\gamma(k) \rightarrow 0$ as $k \rightarrow \infty$.

By a chain of infinite order, we mean a stationary random processes in which, at each step, the probability governing the choice of a new state depends on the entire past. It provides a flexible model that is very useful in diverse areas. For instance, in bioinformatics [3] or linguistics [10]. Chains of infinite order seem to have been first studied by Onicescu and Mihoc [15], who called them chains with complete connections. Their study was soon taken up by Doeblin and Fortes [9] who first proved the results on speed of convergence towards the invariant measure. We refer the reader to Iosifescu and Grigorescu [11] for a complete survey.

A natural approach to studying stationary processes is to approximate the original process by Markov chains of growing order. The conditional probabilities of the canonical approximation of order m coincide with the order m conditional probabilities of the original process. As far as we know, there exists no other results in the literature concerning

the AEP for *non-null* stationary process. It is Wang and Yang's work [17] that will be our setting. This article addresses the following question: How well can we approximate the generalized entropy density of *non-null* stationary stochastic process by a Markov chain of order m ? The significance of this paper is that there are no ergodicity constraints imposed on the process X . We only assume that the process is stationary and *non-null*.

In this paper, first an improvement of a strong limit theorem for the moving averages of the functionals of an m -th order nonhomogeneous Markov chains will be proved by using Borel–Cantelli lemma. Next, as corollaries, some strong limit theorems for the frequencies of occurrence of states in the block $X_{a_n-m+1}^{a_n}, \dots, X_{a_n+\phi(n)-m}^{a_n+\phi(n)-1}$ and the convergence of the generalized relative entropy density for this Markov chains are established. Finally, an explicit bound relating the relative entropy density of the non-null stationary stochastic process and that of the canonical m -order Markov approximation are presented.

Our basic tool is the m -th order canonical Markov approximation technique, which enables us to approximate the *non-null* stationary stochastic process.

We now briefly state our main result and the detailed description can be found in section 3.

Theorem C. *Let $X = (X_n)_{n \in \mathbb{N}}$ be a finite non-null stationary stochastic process with continuity rate $(\beta(k))_{k \in \mathbb{N}}$. If $p_{\inf} > 0$, we have*

$$\begin{aligned} H^{[m]} - \frac{\beta(m)}{p_{\inf}} &\leq \liminf_n f_{a_n, \phi(n)}(\omega) \leq \limsup_n f_{a_n, \phi(n)}(\omega) \\ &\leq H^{[m]} + \frac{\beta(m)}{p_{\inf}} \mathbb{P}|_{\sigma(X)} - a.e. \end{aligned}$$

If X is continuous, then

$$\lim_n f_{a_n, \phi(n)}(\omega) = H^\infty \mathbb{P}|_{\sigma(X)} - a.e.,$$

where $H^{[m]}$ is the entropy of the canonical m -th order Markov approximation of X and $H^\infty = \lim_m H(X_m | X_0^{m-1})$.

The remainder of this paper is organized as follows: Section 2 gives preliminaries in the form of several lemmas. Section 3 is the most important part of the paper, where some limit theorems for m -th order non-homogeneous Markov chains and a new approximation for the relative entropy density of *non-null* stationary process are established. The proofs of the Lemma 3 and Theorem 1 are given in section 4.

2. Some lemmas

We now recall and develop some preliminaries before arriving at the main theorems.

Lemma 1 (Lemma 2 of [17]). *Let $(a_n, \phi(n))_{n \in \mathbb{N}}$ be a sequences of pairs of natural numbers with $\phi(n)$ tending to infinity as $n \rightarrow \infty$. Let $h(x)$ be a bounded function defined on an interval I , and let $(x_n)_{n \in \mathbb{N}}$ be a sequence in I . If*

$$\lim_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} |x_k - x| = 0$$

and $h(x)$ is continuous at point x , then

$$\lim_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} |h(x_k) - h(x)| = 0.$$

Lemma 2 (Lemma 3 of [13]). Let $X = (X_n)_{n \in \mathbb{N}}$ be a stochastic process taking values in finite set \mathbf{X} , and let $f_{a_n, \phi(n)}(\omega)$ be defined by equation (1.2). Then $f_{a_n, \phi(n)}(\omega)$ is uniformly integrable.

Let X be an m -th order nonhomogeneous Markov chain. For $n \geq m$, let

$$\mathbb{P}(X_n = x_n | X_0^{n-1} = x_0^{n-1}) = \mathbb{P}(X_n = x_n | X_{n-m}^{n-1} = x_{n-m}^{n-1}).$$

Set

$$p(i_0^{m-1}) = \mathbb{P}(X_0^{m-1} = i_0^{m-1}),$$

and set

$$p_n(j|i^m) = \mathbb{P}(X_n = j | X_{n-m}^{n-1} = i^m).$$

Here $p(i_0^{m-1})$ is called the m -dimensional initial distribution, $p_n(j|i^m)$ are called the m -th-order transition probabilities and

$$P_n = (p_n(j|i^m)), \quad j \in \mathbf{X}, i^m \in \mathbf{X}^m, \quad n = 1, 2, \dots$$

are called the m -th order transition matrices. In this case,

$$p(x_0^n) = p(x_0^{m-1}) \cdot \prod_{k=m}^n p_k(x_k | x_{k-m}^{k-1}), \quad n \geq m,$$

and the generalized relative entropy density can be written as

$$\begin{aligned} f_{a_n, \phi(n)}^{[m]}(\omega) &= -\frac{1}{\phi(n)} [\log p(X_{a_n+1}^{a_n+\phi(n)})] \\ &= -\frac{1}{\phi(n)} \left\{ \log p(X_{a_n+1}^{a_n+m}) + \sum_{k=a_n+m+1}^{a_n+\phi(n)} \log p_k(X_k | X_{k-m}^{k-1}) \right\}. \end{aligned} \tag{2.1}$$

Lemma 3. Let X be an m -th order nonhomogeneous Markov chain with m -th order initial distribution

$$p(x_0^{m-1}) = \mathbb{P}(X_0^{m-1} = x_0^{m-1}), \quad x_0^{m-1} \in \mathbf{X}^m,$$

and m -th order transition matrices

$$P_n = (p_n(j|i^m)), \quad j \in \mathbf{X}, i^m \in \mathbf{X}^m.$$

Let $(g_n(x^{m+1}))_{n \in \mathbb{N}}$ be a sequence of real functions defined on \mathbf{X}^{m+1} . Suppose $(a_n, \phi(n))_{n \in \mathbb{N}}$ is a sequence of pairs of natural numbers that, for every $\varepsilon > 0$,

$$\sum_{n=1}^{\infty} \exp[-\varepsilon \phi(n)] < \infty. \tag{2.2}$$

If there exists a real number $0 < \gamma < \infty$ such that

$$\limsup_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \mathbb{E}[g_k^2(X_{k-m}^k) e^{\gamma |g_k(X_{k-m}^k)|} |X_{k-m}^{k-1}] = c(\gamma; \omega) < \infty \text{ a.e.}, \quad (2.3)$$

then

$$\lim_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \{g_k(X_{k-m}^k) - \mathbb{E}[g_k(X_{k-m}^k) |X_{k-m}^{k-1}]\} = 0 \text{ a.e.} \quad (2.4)$$

Remark 2. We first note that condition (2.2) can be easily satisfied. For example, let $\phi(n) = [n^\alpha]$ ($\alpha > 0$), where $[\cdot]$ is the usual largest integer function.

Remark 3. Since $\mathbb{E}[g_k(X_{k-m}^k) |X_{k-m}^{k-1}] = \sum_{j=1}^b g_k(X_{k-m}^{k-1}, j) p_k(j |X_{k-m}^{k-1})$, equation (2.4) can be rewritten as

$$\lim_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \left\{ g_k(X_{k-m}^k) - \sum_{j=1}^b g_k(X_{k-m}^{k-1}, j) p_k(j |X_{k-m}^{k-1}) \right\} = 0 \text{ a.e.} \quad (2.5)$$

Remark 4. If $(g_n(x^{m+1}))_{n \in \mathbb{N}}$ are uniformly bounded, then equation (2.3) holds.

By suitable modification to the proof of Lemma 1 in [17], we can give a proof of Lemma 3. For the convenience of readers, we will present the proof in detail in section 4.

COROLLARY 1

Let X be an m -th order nonhomogeneous Markov chain defined as above, and $f_{a_n, \phi(n)}(\omega)$ defined as in equation (2.1). Then

$$\lim_n \left\{ f_{a_n, \phi(n)}^{[m]}(\omega) + \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \sum_{j=1}^b p_k(j |X_{k-m}^{k-1}) \log p_k(j |X_{k-m}^{k-1}) \right\} = 0 \text{ a.e.} \quad (2.6)$$

Let $H(p_1, \dots, p_b)$ be the entropy of the distribution (p_1, \dots, p_b) , i.e.,

$$H(p_1, \dots, p_b) = - \sum_{j=1}^b p_j \log p_j.$$

Equation (2.6) can also be represented as

$$\lim_n \left\{ f_{a_n, \phi(n)}^{[m]}(\omega) - \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} H[p_k(1 |X_{k-m}^{k-1}), \dots, p_k(b |X_{k-m}^{k-1})] \right\} = 0 \text{ a.e.} \quad (2.7)$$

Proof. Putting $g_n(x_{m+1}) = -\log p_n(x_{m+1}|x^m)$ in Lemma 4, by equation (2.5), we have

$$\begin{aligned} & \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \{g_k(X_{k-m}^k) - \sum_{j=1}^b g_k(X_{k-m}^{k-1}, j) \cdot p_k(j|X_{k-m}^{k-1})\} \\ &= -\frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \left\{ \log p_k(X_k|X_{k-m}^{k-1}) - \sum_{j=1}^b p_k(j|X_{k-m}^{k-1}) \log p_k(j|X_{k-m}^{k-1}) \right\} \\ &= \frac{1}{\phi(n)} \log p(X_{a_n-m+1}^{a_n}) + f_{a_n, \phi(n)}(\omega) \\ &+ \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \sum_{j=1}^b p_k(j|X_{k-m}^{k-1}) \log p_k(j|X_{k-m}^{k-1}). \end{aligned}$$

Since

$$\mathbb{E}e^{|\log p(X_{a_n-m+1}^{a_n})|} = \sum_{x_{a_n-m+1}^{a_n} \in \mathbf{X}^m} e^{-\log p(x_{a_n-m+1}^{a_n})} p(x_{a_n-m+1}^{a_n}) = mb,$$

by Markov’s inequality, for any $\varepsilon > 0$, we have

$$\mathbb{P} \left\{ \omega : \frac{1}{\phi(n)} |\log p(X_{a_n-m+1}^{a_n})| \geq \varepsilon \right\} \leq \frac{mb}{e^{\varepsilon\phi(n)}}.$$

Recalling that $\sum_{n=1}^{\infty} \frac{1}{e^{\varepsilon\phi(n)}} < \infty$, we see from Borel–Cantelli lemma that the event

$$\left\{ \omega : \frac{1}{\phi(n)} |\log p(X_{a_n-m+1}^{a_n})| \geq \varepsilon \right\}$$

occurs only finitely often with probability 1. It follows from the arbitrariness of ε that

$$\lim_n \frac{1}{\phi(n)} \log p(X_{a_n-m+1}^{a_n}) \leq 0 \quad \text{a.e.} \tag{2.8}$$

Observe that

$$\mathbb{E}[(\log p_k(X_k|X_{k-m}^{k-1}))^2|X_{k-m}^{k-1}] = \sum_{j=1}^b (\log p_k(j|X_{k-m}^{k-1}))^2 p_k(j|X_{k-m}^{k-1}) \leq 4be^{-2}$$

and that

$$\sum_{k=a_n+1}^{a_n+\phi(n)} \phi(n)^{-1} \mathbb{E}[(\log p_k(X_k|X_{k-m}^{k-1}))^2|X_{k-m}^{k-1}] < \infty. \tag{2.9}$$

Equation (2.6) follows from equations (2.8), (2.9) and Lemma 4. □

Let $N_{a_n, \phi(n)}(i^m; \omega)$ denote the number of occurrences of i^m in the segment sample $X_{a_n - m + 1}^{a_n + \phi(n)}$, i.e.,

$$N_{a_n, \phi(n)}(i^m; \omega) = \text{Card}\{k : X_{k+1}^{k+m} = i^m, a_n - m \leq k \leq a_n + \phi(n) - m\}. \quad (2.10)$$

COROLLARY 2

Let X be an m -th order nonhomogeneous Markov chain defined as in Lemma 4. Then

$$\lim_n \frac{1}{\phi(n)} \left\{ N_{a_n, \phi(n)-1}(i^m; \omega) - \sum_{k=a_n+1}^{a_n+\phi(n)} \mathbf{1}_{\{i^{m-1}\}}(X_{k-m+1}^{k-1}) p_k(i_m | X_{k-m+1}^{k-1}) \right\} = 0 \quad a.e., \quad (2.11)$$

where $\mathbf{1}_A(\cdot)$ is the indicator function of set A .

Proof. Putting $g_k(x^{m+1}) = \mathbf{1}_{\{i^m\}}(x_2^{m+1})$ in Lemma 4, it is not difficult to verify that $\{g_k(x^{m+1})\}_{k=0}^\infty$ satisfies the condition (2.3). Notice that

$$\begin{aligned} & \sum_{k=a_n+1}^{a_n+\phi(n)} \{g_k(X_{k-m}^k) - \sum_{l=1}^b g_k(X_{k-m}^{k-1}, l) p_k(l | X_{k-m}^{k-1})\} \\ &= \sum_{k=a_n+1}^{a_n+\phi(n)} \{\mathbf{1}_{\{i^m\}}(X_{k-m+1}^k) - \sum_{l=1}^b \mathbf{1}_{\{i^{m-1}\}}(X_{k-m+1}^{k-1}) \mathbf{1}_{\{i_m\}}(l) p_k(l | X_{k-m}^{k-1})\} \\ &= N_{a_n, \phi(n)-1}(i^m; \omega) + \mathbf{1}_{\{i^m\}}(X_{a_n+\phi(n)-m+1}^{a_n+\phi(n)}) - \mathbf{1}_{\{i^m\}}(X_{a_n-m+1}^{a_n}) \\ & \quad - \sum_{k=a_n+1}^{a_n+\phi(n)} \mathbf{1}_{\{i^{m-1}\}}(X_{k-m+1}^{k-1}) p_k(i_m | X_{k-m}^{k-1}). \end{aligned} \quad (2.12)$$

Equation (2.11) follows from equation (2.12) and Lemma 3 directly. \square

Let

$$P = (p(j|i^m)), \quad j \in \mathbf{X}, i^m \in \mathbf{X}^m$$

be an m -th order transition matrix. We define a stochastic matrix as follows:

$$\bar{P} = (p(j^m|i^m)), \quad i^m \in \mathbf{X}^m, j^m \in \mathbf{X}^m, \\ p(j^m|i^m) = \begin{cases} p(j_m|i^m), & \text{if } j_k = i_{k+1}, k = 1, 2, \dots, m-1 \\ 0, & \text{otherwise.} \end{cases}$$

\bar{P} is called an m -dimensional stochastic matrix determined by the m -th order transition matrix P .

Lemma 4 (Corollary 2 of [17]). Let \bar{P} be an m -dimensional stochastic matrix determined by the m -th order transition matrix P . If the elements of P are all positive, i.e.,

$$P = (p(j|i^m)), \quad p(j|i^m) > 0, \quad \forall j \in \mathbf{X}, \quad i^m \in \mathbf{X}^m,$$

then \bar{P} is ergodic.

3. Main results

We are now ready to provide the main results of this article.

Theorem 1. Let $X = (X_n)_{n \in \mathbb{N}}$ be an m -th order nonhomogeneous Markov chain defined as in Lemma 3. Let $P = (p(j|i^m))$ be another m -th order transition matrix. Let $(a_n, \phi(n))_{n \in \mathbb{N}}$ and $N_{a_n, \phi(n)}(i^m, \omega)$ be defined as above. Let $f_{a_n, \phi(n)}^{[m]}(\omega)$ be defined as in equation (2.1). Assume that the m -dimensional transition probability matrix \bar{P} determined by P is ergodic. If equation (2.2) holds and

$$\lim_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} |p_k(j|i^m) - p(j|i^m)| = 0, \quad \forall j \in \mathbf{X}, \quad i^m \in \mathbf{X}^m, \quad (3.1)$$

then

$$\lim_n \frac{1}{\phi(n)} N_{a_n, \phi(n)-1}(i^m; \omega) = \pi(i^m) \quad a.e. \quad \forall i^m \in \mathbf{X}^m \quad (3.2)$$

and

$$\lim_n f_{a_n, \phi(n)}^{[m]}(\omega) = - \sum_{i^m \in \mathbf{X}^m} \pi(i^m) \cdot \sum_{j \in \mathbf{X}} p(j|i^m) \log p(j|i^m) \quad a.e., \quad (3.3)$$

where $\{\pi(i^m), i^m \in \mathbf{X}^m\}$ is the unique stationary distribution determined by the transition matrix P .

Remark 5. Putting $a_n = 0$ and $\phi(n) = n$ in Theorem 3, we can obtain the classical Shannon–McMillian–Breiman theorem for m -th order nonhomogeneous Markov chains.

The proof of this theorem will be given in section 4.

COROLLARY 3

Let X be an m -th order homogeneous Markov chain with m -th order transition matrix

$$P = (p(j|i^m)), \quad p(j|i^m) > 0, \quad \forall i^m \in \mathbf{X}^m, \quad j \in \mathbf{X}.$$

Then there exists a distribution

$$\{\pi(i^m), i^m \in \mathbf{X}^m\}$$

such that equations (3.2) and (3.3) hold.

DEFINITION 2

Assume that X is a stationary chain with distribution \mathbb{P} . The canonical m -order Markov approximation of X is the stationary m -order Markov chain (denoted by $X[m]$) compatible with the kernel $P^{[m]}$ defined by (for $n \geq m$)

$$\begin{aligned} P^{[m]}(X_n = i | X_{n-m}^{n-1} = j_{n-m}^{n-1}) &= p^{[m]}(i | j_{n-m}^{n-1}) \\ &= \mathbb{P}(X_n = i | X_{n-m}^{n-1} = j_{n-m}^{n-1}) \quad i \in \mathbf{X}, \quad j_{n-m}^{n-1} \in \mathbf{X}^m. \end{aligned}$$

Set

$$f_{a_n, \phi(n)}^{[m]}(\omega) := -\frac{1}{\phi(n)} \log p^{[m]}(X_{a_n+1}^{a_n+\phi(n)}),$$

where $p^{[m]}(X_{a_n+1}^{a_n+\phi(n)}) = p(X_{a_n+1}^{a_n+m}) \prod_{k=a_n+m+1}^{a_n+\phi(n)} p(X_k | X_{k-m}^{k-1})$.

Theorem 2. *Let $X = (X_n)_{n \in \mathbb{N}}$ be a non-null stationary stochastic process with finitely many values from \mathbf{X} on the probability space $(\Omega, \mathbf{F}, \mathbb{P})$. For each $1 \leq m \leq \phi(n)$, we have*

$$\lim_n f_{a_n, \phi(n)}^{[m]}(\omega) = H^{[m]} \mathbb{P}|_{\sigma(X)} - a.e., \tag{3.4}$$

where $H^{[m]} = -\sum_{i^m \in \mathbf{X}^m} \pi(i^m) \cdot \sum_{j \in \mathbf{X}} p(j|i^m) \log p(j|i^m)$.

Proof. For each $m \geq 1$, if $n \geq m$, let

$$\begin{aligned} p^{[m]}(x_0^n) &= \mathbb{P}(X_0^{m-1} = x_0^{m-1}) \prod_{k=m}^n \mathbb{P}(X_k = x_k | X_{k-m}^{k-1} = x_{k-m}^{k-1}) \\ &= p(x_0^{m-1}) \prod_{k=m}^n p(x_k | x_{k-m}^{k-1}) \end{aligned}$$

and if $0 \leq n < m$, let $p^{[m]}(x_0^n) = \mathbb{P}(X_0^n = x_0^n)$.

The $p^{[m]}$ is a particular Markov measure relevant to \mathbb{P} in the sense that it has the same m -th order transition probabilities as \mathbb{P} . Therefore, by the Kolmogorov’s extension theorem that there exists a probability measure (denoted by $\mathbb{P}^{[m]}$) on (Ω, \mathbf{F}) such that $\mathbb{P}^{[m]}(X_0^n = x_0^n) = p^{[m]}(x_0^n)$, it is easy to show that, under the probability measure $\mathbb{P}^{[m]}$, X is an m -th order stationary homogeneous Markov chain with positive transition matrix

$$P = (p(j|i^m)), \quad j \in \mathbf{X}, \quad i^m \in \mathbf{X}^m.$$

Since $p(j|i^m) > 0$, $j \in \mathbf{X}$, $i^m \in \mathbf{X}^m$, by Corollary 3, we have

$$\lim_n f_{a_n, \phi(n)}^{[m]}(\omega) = \lim_n -\frac{1}{\phi(n)} \log p^{[m]}(X_{a_n+1}^{a_n+\phi(n)}) = \text{a constant } \mathbb{P}^{[m]} - a.e. \tag{3.5}$$

Note that

$$\begin{aligned} & \mathbb{E}_{\mathbb{P}^{[m]}} f_{a_n, \phi(n)}^{[m]}(\omega) \\ &= \mathbb{E}_{\mathbb{P}^{[m]}} \left\{ -\frac{1}{\phi(n)} \log p^{[m]}(X_{a_n+m}^{a_n+\phi(n)}) - \frac{1}{\phi(n)} \sum_{k=m}^{a_n+\phi(n)} \log p(X_k | X_{k-m}^{k-1}) \right\} \\ &= \frac{\mathbb{E}_{\mathbb{P}^{[m]}} \{-\log p(X_0^m)\}}{\phi(n)} + \frac{\phi(n) - m}{\phi(n)} \mathbb{E}_{\mathbb{P}^{[m]}} \{-\log p(X_0 | X_m^{-1})\} \text{ (by stationarity)} \\ &\rightarrow \mathbb{E}_{\mathbb{P}^{[m]}} \{-\log p(X_0 | X_m^{-1})\} \text{ as } n \rightarrow \infty, \end{aligned}$$

where $\mathbb{E}_{\mathbb{P}^{[m]}}$ denotes taking expectation under the probability measure $\mathbb{P}^{[m]}$.

From Lemma 2, $f_{a_n, \phi(n)}^{[m]}(\omega)$ is uniformly integrable under the measure $\mathbb{P}^{[m]}$, we have

$$\lim_n \int_{\Omega} f_{a_n, \phi(n)}^{[m]}(\omega) d\mathbb{P}^{[m]} = \lim_n \mathbb{E}_{\mathbb{P}^{[m]}} f_{a_n, \phi(n)}^{[m]}(\omega) = \text{the constant.}$$

Therefore, the constant in equation (3.5) is equal to $\mathbb{E}_{\mathbb{P}^{[m]}} \{-\log p(X_0 | X_m^{-1})\}$, i.e.,

$$\lim_n f_{a_n, \phi(n)}^{[m]}(\omega) = \mathbb{E}_{\mathbb{P}^{[m]}} \{-\log p(X_0 | X_m^{-1})\} \quad \mathbb{P}^{[m]} - \text{a.e.} \tag{3.6}$$

Restricting the measure \mathbb{P} to the trajectory space of X (denoted by $\mathbb{P}|_{\sigma(X)}$), it is not difficult to verify that $\mathbb{P}|_{\sigma(X)} \ll \mathbb{P}^{[m]}$, therefore, we have by equations (3.5) and (3.6) and the fact that $\mathbb{E}_{\mathbb{P}^{[m]}} \{-\log p(X_0 | X_m^{-1})\} = \mathbb{E} \{-\log p(X_0 | X_m^{-1})\} = H^{[m]}$, that

$$\lim_n f_{a_n, \phi(n)}^{[m]}(\omega) = \mathbb{E} \{-\log p(X_m | X_0^{m-1})\} = H^{[m]} \quad \mathbb{P}|_{\sigma(X)} - \text{a.e.,}$$

which concludes the proof of the theorem. □

We remark that the measures \mathbb{P} and $\mathbb{P}^{[m]}$ cannot be compared with each other.

In classical information theory, the following equation

$$\lim_n -\frac{1}{n} \log p(X^n) = \text{a constant a.e.} \tag{3.7}$$

holds for finite stationary ergodic sequences of random variables, which is the famous Shannon–MacMillan theorem. A natural problem is whether the equation also holds for *non-null* stationary process? The following two examples show that the notations of *non-null* and ergodicity do not coincide, i.e., a stationary ergodic sequence of random variables may not be *non-null* and a *non-null* stationary sequence of random variables may not be ergodic and, unfortunately, equation (3.7) does not hold for non-null stationary process.

Example 1. Let $X^{(1)} = (X_n^{(1)})_{n \in \mathbb{N}^*}$ and $X^{(2)} = (X_n^{(2)})_{n \in \mathbb{N}^*}$ be two non-null stationary ergodic processes with values in \mathbf{X} . By the Shannon–McMillan–Breiman theorem [1],

$$\begin{aligned} \lim_n -\frac{1}{n} \log p(X_1^{(1)}, \dots, X_n^{(1)}) &= H_1 \quad \text{a.e.,} \\ \lim_n -\frac{1}{n} \log p(X_1^{(2)}, \dots, X_n^{(2)}) &= H_2 \quad \text{a.e.,} \end{aligned}$$

where H_1, H_2 are the entropies of $X^{(1)}$ and $X^{(2)}$ respectively. Assume that $H_1 \neq H_2$. Suppose $A \in \mathbf{F}$ with $0 < \mathbb{P}(A) < 1$ and suppose that A is independent of the processes $X^{(1)}$ and $X^{(2)}$. Define a new process $X^{(3)} = (X_n^{(3)})_{n \in \mathbb{N}^*}$ on $(\Omega, \mathbf{F}, \mathbb{P})$ as follows: If $\omega \in A$, let $X_n^{(3)} = X_n^{(1)}$ for all $n \in \mathbb{N}^*$, and if $\omega \in A^c$, let $X_n^{(3)} = X_n^{(2)}$ for all $n \in \mathbb{N}^*$. It is obvious that the set $\{A, A^c\}$ is invariant. Next, we shall show that the process $X^{(3)}$ defined above is non-null and stationary but not ergodic.

Note that for any $k, n \geq 1, x^n \in \mathbf{X}^n$. Then

$$\begin{aligned}
 & \mathbb{P}\{(X_1^{(3)}, \dots, X_n^{(3)}) = x^n\} \\
 &= \mathbb{P}(A)\mathbb{P}\{(X_1^{(3)}, \dots, X_n^{(3)}) = x^n | A\} + \mathbb{P}(A^c)\mathbb{P}\{(X_1^{(3)}, \dots, X_n^{(3)}) = x^n | A^c\} \\
 &= \mathbb{P}(A)\mathbb{P}\{(X_1^{(1)}, \dots, X_n^{(1)}) = x^n | A\} + \mathbb{P}(A^c)\mathbb{P}\{(X_1^{(2)}, \dots, X_n^{(2)}) = x^n | A^c\} \\
 &= \mathbb{P}(A)\mathbb{P}\{(X_1^{(1)}, \dots, X_n^{(1)}) = x^n\} + \mathbb{P}(A^c)\mathbb{P}\{(X_1^{(2)}, \dots, X_n^{(2)}) = x^n\} \quad (3.8) \\
 &= \mathbb{P}(A)\mathbb{P}\{(X_{1+k}^{(1)}, \dots, X_{n+k}^{(1)}) = x^n\} + \mathbb{P}(A^c)\mathbb{P}\{(X_{1+k}^{(2)}, \dots, X_{n+k}^{(2)}) = x^n\} \\
 &= \mathbb{P}(A)\mathbb{P}\{(X_{1+k}^{(1)}, \dots, X_{n+k}^{(1)}) = x^n | A\} + \mathbb{P}(A^c)\mathbb{P}\{(X_{1+k}^{(2)}, \dots, X_{n+k}^{(2)}) = x^n | A^c\} \\
 &= \mathbb{P}(A)\mathbb{P}\{(X_{1+k}^{(3)}, \dots, X_{n+k}^{(3)}) = x^n | A\} + \mathbb{P}(A^c)\mathbb{P}\{(X_{1+k}^{(3)}, \dots, X_{n+k}^{(3)}) = x^n | A^c\} \\
 &= \mathbb{P}\{(X_{1+k}^{(3)}, \dots, X_{n+k}^{(3)}) = x^n\}. \quad (3.9)
 \end{aligned}$$

It is easy to see from equations (3.8) and (3.9) that the process $X^{(3)}$ is *non-null* and stationary and that

$$\begin{aligned}
 \lim_n -\frac{1}{n} \log p(X_1^{(3)}, \dots, X_n^{(3)}) &= H_1 \quad \text{a.e. } \omega \in A, \\
 \lim_n -\frac{1}{n} \log p(X_1^{(3)}, \dots, X_n^{(3)}) &= H_2 \quad \text{a.e. } \omega \in A^c.
 \end{aligned}$$

Notice that $H_1 \neq H_2$, and hence $X^{(3)}$ cannot be ergodic.

Example 2. Consider a homogeneous Markov chain $X = (X_n)_{n \in \mathbb{N}^*}$ with state space $\{1, 2, 3\}$ and transition matrix

$$P = \begin{pmatrix} 0 & 2/3 & 1/3 \\ 1/3 & 0 & 2/3 \\ 2/3 & 1/3 & 0 \end{pmatrix}.$$

It is not difficult to check that the unique invariant probability of the chain is $\pi = (1/3, 1/3, 1/3)$, hence it is ergodic, but $\mathbb{P}(X_1 = 1, X_2 = 1) = 0$.

Example 1 indicates that under the assumption of being *non-null* and stationary can not guarantee the existence of $\lim_n -\frac{1}{n} \log p(X^n)$. In the following Theorem 3, we will try to fill this gap to some extent. We give the upper and lower bounds of $f_{a_n, \phi(n)}(\omega)$ expressed using the concepts of continuous rate and log-continuous rate. At the same time, under some mild assumptions, we establish a weak form of the generalized ergodic theorem.

Theorem 3. Let $X = (X_n)_{n \in \mathbb{N}}$ be a finite non-null stationary stochastic process with continuity rate $(\beta(k))_{k \in \mathbb{N}}$ and $X[m]$ be the canonical m -order Markov approximation of

the process. Under the strong non-nullness assumption, $\inf_{n,x_0^{n-1}} p(i|x_0^{n-1}) \geq p_{\text{inf}} > 0$ for $\phi(n) > m$, and we have

$$\begin{aligned} H^{[m]} - \frac{\beta(m)}{p_{\text{inf}}} &\leq \liminf_n f_{a_n, \phi(n)}(\omega) \leq \limsup_n f_{a_n, \phi(n)}(\omega) \\ &\leq H^{[m]} + \frac{\beta(m)}{p_{\text{inf}}} \mathbb{P}|_{\sigma(X)} - a.e. \end{aligned} \tag{3.10}$$

Furthermore, if X is continuous, then

$$\lim_n f_{a_n, \phi(n)}(\omega) = H^\infty \mathbb{P}|_{\sigma(X)} - a.e., \tag{3.11}$$

where $H^{[m]} = -\sum_{i^m \in \mathbf{X}^m} \pi(i^m) \cdot \sum_{j \in \mathbf{X}} p(j|i^m) \log p(j|i^m)$ and $H^\infty = \lim_m H(X_m|X_0^{m-1})$.

Proof. Applying the inequality $\log x \leq x - 1$ ($x > 0$) and equation (1.4), we have for $\phi(n) > m$,

$$\begin{aligned} &\frac{1}{\phi(n)} |\log p(x_{a_n+\phi(n)}^{a_n+\phi(n)}) - \log p^{[m]}(x_{a_n+1}^{a_n+\phi(n)})| \\ &= \frac{1}{\phi(n)} \left| \log p(x_{a_n+1}^{a_n+m}) \prod_{k=a_n+m+1}^{a_n+\phi(n)} p(x_k|x_{a_n+1}^{k-1}) - \log p(x_{a_n+1}^{a_n+m}) \prod_{k=a_n+m+1}^{a_n+\phi(n)} p(x_k|x_{k-m}^{k-1}) \right| \\ &\leq \frac{1}{\phi(n)} \{ |\log[p(x_{a_n+m+2}|X_{a_n+1}^{a_n+m+1})/p(x_{a_n+m+2}|x_{a_n+2}^{a_n+m+1})]| \\ &\quad + |\log[p(x_{a_n+m+3}|x_{a_n+1}^{a_n+m+2})/p(x_{a_n+m+3}|x_{a_n+3}^{a_n+m+2})]| \\ &\quad + \dots + |\log[p(x_{a_n+\phi(n)}|x_{a_n+1}^{a_n+\phi(n)-1})/p(x_{a_n+\phi(n)}|x_{a_n+\phi(n)-m}^{a_n+\phi(n)-1})]| \} \\ &\leq \frac{1}{\phi(n)} \left[\frac{|p(x_{a_n+m+2}|x_{a_n+1}^{a_n+m+1}) - p(x_{a_n+m+2}|x_{a_n+2}^{a_n+m+1})|}{p(x_{a_n+m+2}|x_{a_n+1}^{a_n+m+1})} \right. \\ &\quad \left. + \dots + \frac{|p(x_{a_n+\phi(n)}|x_{a_n+1}^{a_n+\phi(n)-1}) - p(x_{a_n+\phi(n)}|x_{a_n+\phi(n)-m}^{a_n+\phi(n)-1})|}{p(x_{a_n+\phi(n)}|x_{a_n+1}^{a_n+\phi(n)-1})} \right] \\ &= \frac{1}{\phi(n)} \\ &\quad \times \left[\frac{|P(X_0 = x_{a_n+m+2}|X_{-m-1}^{-1} = x_{a_n+1}^{a_n+m+1}) - P(X_0 = x_{a_n+m+2}|X_{-m}^{-1} = x_{a_n+2}^{a_n+m+1})|}{P(X_0 = x_{a_n+m+2}|X_{-m-1}^{-1} = x_{a_n+1}^{a_n+m+1})} \right. \\ &\quad \left. + \dots + \frac{|P(X_0 = x_{a_n+\phi(n)}|X_{-\phi(n)-1}^{-1} = x_{a_n+1}^{a_n+\phi(n)-1}) - P(X_0 = x_{a_n+\phi(n)}|X_{-m}^{-1} = x_{a_n+\phi(n)-m}^{a_n+\phi(n)-1})|}{P(X_0 = x_{a_n+\phi(n)}|X_{-\phi(n)-1}^{-1} = x_{a_n+1}^{a_n+\phi(n)-1})} \right] \end{aligned}$$

$$\begin{aligned} &\leq \frac{1}{\phi(n)} \left[\frac{\beta(m)}{p_{\inf}} + \cdots + \frac{\beta(m)}{p_{\inf}} \right] \text{ (by equation (1.3) and the stationarity)} \\ &= \frac{\beta(m)}{p_{\inf}}, \end{aligned}$$

that is,

$$|f_{a_n, \phi(n)}(\omega) - f_{a_n, \phi(n)}^{[m]}(\omega)| \leq \frac{\beta(m)}{p_{\inf}}. \quad (3.12)$$

Thus, based on the above bound for finite samples of size n , equation (3.10) follows immediately from equation (3.6).

It is well known that $\lim_m H(X_m | X_0^{m-1})$ always exists (denoted by H^∞) for finite stationary processes. Let $m \rightarrow +\infty$ on both sides of equation (3.10), then equation (3.11) follows. \square

Remark 6. It is easy to see that if the continuity rate $(\beta(k))_{k \in \mathbb{N}}$ in Theorem 2 is substituted by log-continuity rate $(\gamma(k))_{k \in \mathbb{N}}$, then we have

$$\begin{aligned} H^{[m]} - \gamma(m) &\leq \liminf_n f_{a_n, \phi(n)}(\omega) \leq \limsup_n f_{a_n, \phi(n)}(\omega) \\ &\leq H^{[m]} + \gamma(m) \quad \mathbb{P}|_{\sigma(X)} - \text{a.e.} \end{aligned} \quad (3.13)$$

Moreover, if X is log-continuous, then equation (3.11) also holds.

In this paper, we consider statistical estimates based on a sample $X_{a_n+1}^{a_n+\phi(n)}$ of length $\phi(n)$ of the process. For $\phi(n) \geq m$, the generalized empirical probability of the string i^m is

$$\hat{\pi}(i^m) := \frac{N_{a_n, \phi(n)}(i^m)}{\phi(n)},$$

where $N_{a_n, \phi(n)}(i^m)$ is defined as in equation (2.10). The generalized empirical conditional probability of $j \in \mathbf{X}$ given by i^m is

$$\hat{p}(j|i^m) := \frac{N_{a_n, \phi(n)}(i^m j)}{N_{a_n, \phi(n)-1}(i^m)}.$$

Replacing in equation (3.3) the probabilities by their estimators, we get the following estimator of m -order blockwise empirical entropy

$$\hat{H}_{a_n, \phi(n)}^{[m]}(\omega) := -\frac{1}{\phi(n)} \sum_{i^m \in \mathbf{X}^m} \hat{\pi}(i^m) \sum_{j \in \mathbf{X}} \hat{p}(j|i^m) \log \hat{p}(j|i^m).$$

4. The proofs

Proof of Lemma 3. Let s be a nonzero real number and define

$$\Lambda_{a_n, \phi(n)}(s, \omega) = \frac{\exp\{s \sum_{k=a_n+1}^{a_n+\phi(n)} g_k(X_{k-m}^k)\}}{\prod_{k=a_n+1}^{a_n+\phi(n)} \mathbb{E}[e^{s g_k(X_{k-m}^k)} | X_{k-m}^{k-1}]}, \quad n = 1, 2, \dots,$$

and note that

$$\begin{aligned} & \mathbb{E}\Lambda_{a_n, \phi(n)}(s, \omega) \\ &= \mathbb{E}\{\mathbb{E}[\Lambda_{a_n, \phi(n)}(s, \omega) | X_0^{a_n + \phi(n) - 1}]\} \\ &= \mathbb{E}\left\{ \mathbb{E}\left[\Lambda_{a_n, \phi(n) - 1}(s, \omega) \cdot \frac{e^{s g_{a_n + \phi(n)}(X_{a_n + \phi(n) - m}^{a_n + \phi(n)})}}{\mathbb{E}[e^{s g_{a_n + \phi(n)}(X_{a_n + \phi(n) - m}^{a_n + \phi(n)}) | X_{a_n + \phi(n) - m}^{a_n + \phi(n) - 1}]} \mid X_0^{a_n + \phi(n) - 1} \right] \right\} \\ &= \mathbb{E}\left[\frac{\Lambda_{a_n, \phi(n) - 1}(s, \omega) \mathbb{E}[e^{s g_{a_n + \phi(n)}(X_{a_n + \phi(n) - m}^{a_n + \phi(n)}) | X_{a_n + \phi(n) - m}^{a_n + \phi(n) - 1}]}{\mathbb{E}[e^{s g_{a_n + \phi(n)}(X_{a_n + \phi(n) - m}^{a_n + \phi(n)}) | X_{a_n + \phi(n) - m}^{a_n + \phi(n) - 1}]} \right] \text{ (by Markov property)} \\ &= \mathbb{E}\Lambda_{a_n, \phi(n) - 1}(s, \omega) = \dots = \mathbb{E}\Lambda_{a_n, 1}(s, \omega) = 1. \end{aligned}$$

By a similar argument of equation (2.8), we get

$$\limsup_n \frac{1}{\phi(n)} \log \Lambda_{a_n, \phi(n)}(s, \omega) \leq 0 \quad \text{a.e.} \tag{4.1}$$

Note that

$$\begin{aligned} & \frac{1}{\phi(n)} \log \Lambda_{a_n, \phi(n)}(s, \omega) \\ &= \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n + \phi(n)} \{s g_k(X_{k-m}^k) - \log \mathbb{E}[e^{s g_k(X_{k-m}^k) | X_{k-m}^{k-1}]\}. \end{aligned} \tag{4.2}$$

Equations (4.1) and (4.2) imply that

$$\limsup_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n + \phi(n)} \{s g_k(X_{k-m}^k) - \log \mathbb{E}[e^{s g_k(X_{k-m}^k) | X_{k-m}^{k-1}]\} \leq 0 \quad \text{a.e.} \tag{4.3}$$

Letting $0 < s < \gamma$ and dividing both sides of equation (4.3) by s , we obtain that

$$\limsup_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n + \phi(n)} \{g_k(X_{k-m}^k) - \frac{1}{s} \log \mathbb{E}[e^{s g_k(X_{k-m}^k) | X_{k-m}^{k-1}]\} \leq 0 \quad \text{a.e.} \tag{4.4}$$

Using the elementary inequalities $\log x \leq x - 1$ ($x > 0$) and $0 \leq e^x - 1 - x \leq \frac{1}{2}x^2e^{|x|}$ ($x \in \mathbb{R}$), by equation (4.4), we obtain that

$$\begin{aligned} & \limsup_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n + \phi(n)} \{g_k(X_{k-m}^k) - \mathbb{E}[g_k(X_{k-m}^k) | X_{k-m}^{k-1}]\} \\ & \leq \limsup_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n + \phi(n)} \left\{ \frac{1}{s} \log \mathbb{E}[e^{s g_k(X_{k-m}^k) | X_{k-m}^{k-1}}] - \mathbb{E}[g_k(X_{k-m}^k) | X_{k-m}^{k-1}] \right\} \end{aligned}$$

$$\begin{aligned}
 &\leq \limsup_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \left\{ \frac{\mathbb{E}[(e^{sg_k(X_{k-m}^k)} - 1 - sg_k(X_{k-m}^k)) | X_{k-m}^{k-1}]}{s} \right\} \\
 &\leq \frac{s}{2} \limsup_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \mathbb{E}[g_k^2(X_{k-m}^k) e^{s|g_k(X_{k-m}^k)|} | X_{k-m}^{k-1}] \\
 &\leq \frac{1}{2} sc(\gamma; \omega) < \infty \quad \text{a.e.}
 \end{aligned} \tag{4.5}$$

Letting $s \downarrow 0^+$ in equation (4.5), we have

$$\limsup_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} [g_k(X_{k-m}^k) - \mathbb{E}(g_k(X_{k-m}^k) | X_{k-m}^{k-1})] \leq 0 \quad \text{a.e.} \tag{4.6}$$

Letting $-\gamma < s < 0$ in equation (4.3), and proceeding as in the proof of equation (4.6), we have that

$$\liminf_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} [g_k(X_{k-m}^k) - \mathbb{E}(g_k(X_{k-m}^k) | X_{k-m}^{k-1})] \geq 0 \quad \text{a.e.} \tag{4.7}$$

Equation (2.4) now follows immediately from equations (4.6) and (4.7). □

Proof of Theorem 1. From Corollary 2, we have that

$$\begin{aligned}
 &\lim_n \frac{1}{\phi(n)} \left\{ N_{a_n, \phi(n)-1}(i^m) - \sum_{k=a_n+1}^{a_n+\phi(n)} \sum_{l=1}^b \mathbf{1}_{\{l\}}(X_{k-m}) \mathbf{1}_{\{i^{m-1}\}}(X_{k-m+1}^{k-1}) p_k(i_m | l, i^{m-1}) \right\} \\
 &= 0 \quad \text{a.e.}
 \end{aligned} \tag{4.8}$$

It is not difficult to see from equation (3.1) that

$$\begin{aligned}
 &\lim_n \left| \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \sum_{l=1}^b \mathbf{1}_{\{l\}}(X_{k-m}) \mathbf{1}_{\{i^{m-1}\}}(X_{k-m+1}^{k-1}) [p_k(j | l, i^{m-1}) - p(j | l, i^{m-1})] \right| \\
 &\leq \sum_{l=1}^b \lim_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} |p_k(j | l, i^{m-1}) - p(j | l, i^{m-1})| = 0, \quad \forall j \in \mathbf{X}.
 \end{aligned} \tag{4.9}$$

Combining equations (4.8) and (4.9), we obtain

$$\lim_n \frac{1}{\phi(n)} \left[N_{a_n, \phi(n)-1}(i^m; \omega) - \sum_{l=1}^b N_{a_n, \phi(n)-1}(l, i^{m-1}; \omega) p(i_m | l, i^{m-1}) \right]$$

$$\begin{aligned}
 &= \lim_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \sum_{l=1}^b \mathbf{1}_{\{l\}}(X_{k-m}) [p_k(i_m|l, i^{m-1}) - p(i_m|l, i^{m-1})] \\
 &= 0, \quad \text{a.e.}
 \end{aligned}
 \tag{4.10}$$

Set $k^m = (k_1, \dots, k_m)$, by equation (4.10), we have

$$\lim_n \left\{ \frac{N_{a_n, \phi(n)-1}(i^m)}{\phi(n)} - \sum_{k^m \in \mathbf{X}^m} \frac{N_{a_n, \phi(n)-1}(k^m) - 1}{\phi(n)} p(i^m|k^m) \right\} = 0 \text{ a.e.}$$

(4.11)

Multiplying both sides of equation (4.11) by $p(j^m|i^m)$ and adding them together for $j^m \in \mathbf{X}^m$, we have by equation (4.7) that

$$\begin{aligned}
 0 &= \lim_n \frac{1}{\phi(n)} \left[\sum_{i^m \in \mathbf{X}^m} N_{a_n, \phi(n)-1}(i^m; \omega) p(j^m|i^m) \right. \\
 &\quad \left. - \sum_{i^m \in \mathbf{X}^m} \sum_{k^m \in \mathbf{X}^m} N_{a_n, \phi(n)-1}(k^m; \omega) p(i^m|k^m) p(j^m|i^m) \right] \\
 &= \lim_n \left[\sum_{i^m \in \mathbf{X}^m} \frac{1}{\phi(n)} N_{a_n, \phi(n)-1}(i^m; \omega) p(j^m|i^m) - \frac{1}{\phi(n)} N_{a_n, \phi(n)-1}(j^m; \omega) \right] \\
 &\quad + \lim_n \left[\frac{1}{\phi(n)} N_{a_n, \phi(n)-1}(j^m; \omega) \right. \\
 &\quad \left. - \sum_{k^m \in \mathbf{X}^m} \sum_{i^m \in \mathbf{X}^m} \frac{1}{\phi(n)} N_{a_n, \phi(n)-1}(k^m; \omega) p(j^m|i^m) p(i^m|k^m) \right] \\
 &= \lim_n \left[\frac{1}{\phi(n)} N_{a_n, \phi(n)-1}(j^m; \omega) - \sum_{k^m \in \mathbf{X}^m} \frac{1}{\phi(n)} N_{a_n, \phi(n)-1}(k^m; \omega) p^{(2)}(j^m|k^m) \right] \text{ a.e.,}
 \end{aligned}$$

(4.12)

where $p^{(l)}(j^m|i^m)$ (l is a positive integer) is the l -step transition probability determined by the transition matrix \bar{P} . By induction, we have for any positive integer h that

$$\lim_n \frac{1}{\phi(n)} \left[N_{a_n, \phi(n)-1}(j^m; \omega) - \sum_{k_1^m \in \mathbf{X}^m} N_{a_n, \phi(n)-1}(k_1^m; \omega) p^{(h)}(j^m|k_1^m) \right] = 0, \text{ a.e.}$$

(4.13)

It is easy to see that $\sum_{k^m \in \mathbf{X}^m} N_{a_n, \phi(n)-1}(k^m, \omega) = \phi(n)$. Since \bar{P} is ergodic, we have

$$\lim_h p^{(h)}(j^m|k^m) = \pi(j^m), \quad \forall k^m \in \mathbf{X}^m.$$

(4.14)

Equation (3.2) follows from equations (4.13) and (4.14).

If equation (3.1) holds, it is easy to see from Lemma 2 that

$$\lim_n \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} |p_k(j|i^m) \log p_k(j|i^m) - p(j|i^m) \log p(j|i^m)| = 0, \quad \forall j \in \mathbf{X}, i^m \in \mathbf{X}^m. \tag{4.15}$$

Notice that

$$\begin{aligned} & \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \sum_{j=1}^b p_k(j|X_{k-m}^{k-1}) \log p_k(j|X_{k-m}^{k-1}) \\ &= \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \sum_{j=1}^b \sum_{i^m \in \mathbf{X}^m} \mathbf{1}_{\{i^m\}}(X_{k-m}^{k-1}) \cdot p_k(j|i^m) \log p_k(j|i^m) \end{aligned}$$

implies that

$$\begin{aligned} & \left| f_{a_n, \phi(n)}^{[m]}(\omega) + \sum_{i^m \in \mathbf{X}^m} \pi(i^m) \sum_{j=1}^b p(j|i^m) \log p(j|i^m) \right| \\ & \leq \left| f_{a_n, \phi(n)}^{[m]}(\omega) + \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \sum_{j=1}^b \sum_{i^m \in \mathbf{X}^m} \mathbf{1}_{\{i^m\}}(X_{k-m}^{k-1}) \cdot p_k(j|i^m) \log p_k(j|i^m) \right| \\ & \quad + \left| \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \sum_{j=1}^b \sum_{i^m \in \mathbf{X}^m} \mathbf{1}_{\{i^m\}}(X_{k-m}^{k-1}) \cdot [p_k(j|i^m) \log p_k(j|i^m) - p(j|i^m) \log p(j|i^m)] \right| \\ & \quad + \left| \sum_{i^m \in \mathbf{X}^m} \left[\frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \mathbf{1}_{\{i^m\}}(X_{k-m}^{k-1}) - \pi(i^m) \right] \cdot \sum_{j=1}^b p(j|i^m) \log p(j|i^m) \right| \\ & \leq \left| f_{a_n, \phi(n)}^{[m]}(\omega) + \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \sum_{j=1}^b p_k(j|X_{k-m}^{k-1}) \right| \\ & \quad + \sum_{j=1}^b \frac{1}{\phi(n)} \sum_{k=a_n+1}^{a_n+\phi(n)} \sum_{i^m \in \mathbf{X}^m} |p_k(j|i^m) \log p_k(j|i^m) - p(j|i^m) \log p(j|i^m)| \\ & \quad + \sum_{i^m \in \mathbf{X}^m} \left| \frac{N_{a_n, \phi(n)-1}(i^m)}{\phi(n)} - \pi(i^m) \right| \cdot \left| \sum_{j=1}^b p(j|i^m) \log p(j|i^m) \right|. \tag{4.16} \end{aligned}$$

Equation (3.3) now follows from equations (2.1), (3.2), (4.15) and (4.16) as required. \square

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References

- [1] Algoet P H and Cover T M, A sandwich proof of the Shannon–McMillan–Breiman theorem, *Ann. Probab.* **16** (1988) 899–909
- [2] Barron A R, The strong ergodic theorem for densities: generalized Shannon–McMillan–Breiman theorem, *Ann. Probab.* **13** (1985) 1292–1303
- [3] Bejerano G and Yona G, Variations on probabilistic suffix trees: statistical modeling and prediction of protein families, *Bioinformatics* **17**(1) (2001) 23–43
- [4] Billingsley P, Ergodic theory and information (1965) (New York: Wiley)
- [5] Breiman L, The individual ergodic theorem of information theory, *Ann. Math. Statist.* **28** (1957) 809–811
- [6] Breiman L, A correction to ‘The individual ergodic theorem of information theory’, *Ann. Math. Statist.* **31** (1960) 809–810
- [7] Chung K L, The ergodic the theorem of information theory, *Ann. Math. Statist.* **32** (1961) 612–614
- [8] Csiszár I and Talata Z, On rate of convergence of statistical estimation of stationary ergodic processes, *IEEE Trans. Inform. Theory* **56**(8) (2010) 3637–3641
- [9] Doebelin W and Fortet R, Sur les chaînes à liaisons complètes, *Bull. Soc. Math. France* **65** (1937) 132–148
- [10] Galves A and Leonardi F, Context tree selection and linguistic rhythm retrieval from written texts, *Ann. Appl. Statist.* **6**(1) (2012) 186–209
- [11] Iosifescu M and Grigorescu S, Dependence with complete connections and its applications (1990) (Cambridge: Cambridge University Press)
- [12] Kiefer J C, A simple proof of the Moy–Perez generalization of the Shannon–McMillan theorem, *Pacific J. Math.* **51** (1974) 203–204
- [13] McMillan B, The basic theorem of information theory, *Ann. Math. Statist.* **24** (1953) 196–215
- [14] Nair R, On moving averages and asymptotic equipartition of information, *Period. Math. Hungar.* **71** (2015) 59–63
- [15] Onicescu O and Mihoc G, Sur les chaînes statistiques, *C. R. Acad. Sci. Paris* **200** (1935) 511–522
- [16] Shannon C A, A mathematical theorem of communication, *Bell System Teach. J.* **27** (1948) 379–423, 623–656
- [17] Wang Z Z and Yang W G, The generalized entropy ergodic theorem for nonhomogeneous Markov chains, *J. Theor. Probab.* **29** (2016) 761–775
- [18] Yang W G, The asymptotic equipartition property for nonhomogeneous Markov information sources, *Probab. Eng. Inform. Sc.* **12** (1998) 509–518
- [19] Yang W G and Liu W, The asymptotic equipartition property for M th-order nonhomogeneous Markov information sources, *IEEE Trans. Inform. Theory* **50**(12) (2004) 3326–3330
- [20] Zhong P P, Yang W G and Liang P P, The asymptotic equipartition property for asymptotic circular Markov chains, *Probab. Eng. Inform. Sc.* **24** (2010) 279–288